

## **CSC535: Probabilistic Graphical Models**

#### **Probability Primer : Continuous Probability**

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#### Administrative Items

# Do not use Gradescope! Use D2L!

- If it's not on D2L it won't get graded and you'll receive a zero
- This is the 3<sup>rd</sup> announcement
- HW1 Due Tonight @ 11:59 PM (**D2L**)
- HW2 Out Tonight (will announce on Piazza)

## Outline

#### Motivation and Definitions

- Fundamental Rules of Probability (recap)
- Useful Continuous Distributions / Densities

## Outline

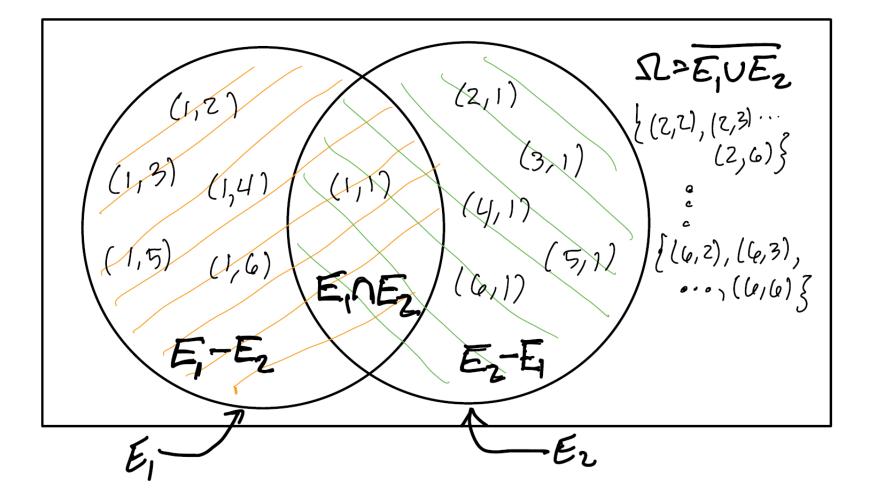
#### Motivation and Definitions

Fundamental Rules of Probability (recap)

> Useful Continuous Distributions / Densities

## **Probability Space**

Recall that we can think of outcomes of a random experiment as a space...



...this has a formal mathematical definition

## **Probability Space**

A sample space  $\Omega$ : set of all possible outcomes of the experiment.

An **event space**  $\mathcal{F}$ : Sets of <u>allowable events</u>  $E \in \mathcal{F}$ 

A probability function  $P: \mathcal{F} \to \mathbf{R}$  satisfying:

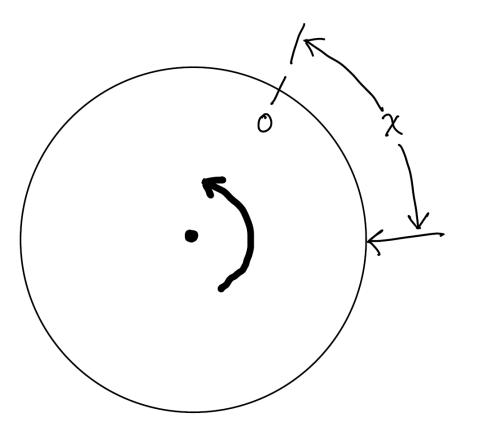
1. For any event  $E, 0 \le P(E) \le 1$ 2.  $P(\Omega) = 1$  and  $P(\emptyset) = 0$ 

Axioms of Probability

3. For any *finite* or *countably infinite* sequence of pairwise mutually disjoint events  $E_1, E_2, E_3, \ldots$ 

$$P\Big(\bigcup_{i\geq 1} E_i\Big) = \sum_{i\geq 1} P(E_i)$$

**Experiment** Spin continuous wheel and measure X displacement from 0



**Question** Assuming uniform probability, what is p(X = x)?

 $\blacktriangleright$  Let  $p(X = x) = \pi$  be the probability of any single outcome

► Let S(k) be set of any k *distinct* points in [0, 1) then,  $P(x \in S(k)) = k\pi$ 

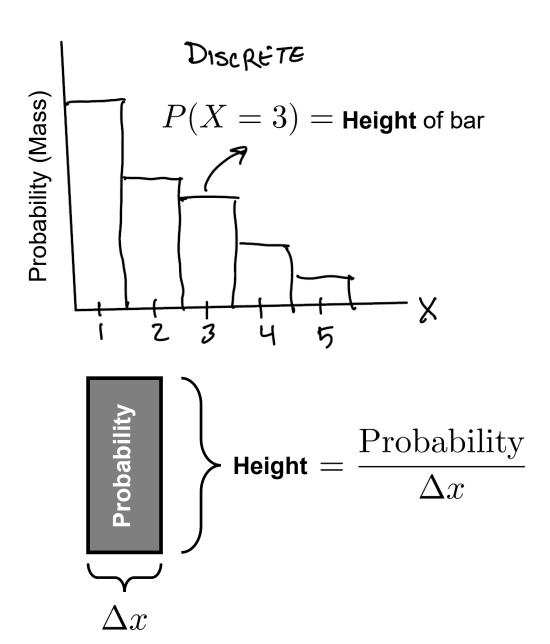
Since  $0 < P(x \in S(k)) < 1$  by axioms of probability,  $k\pi < 1$  for any k

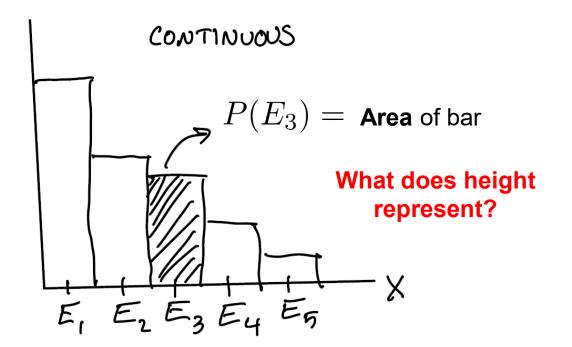
For the refore:  $\pi = 0$  and  $P(x \in S(k)) = p(X = x) = 0$ 

- $\succ$  We have a well-defined event that x takes a value in set  $x \in S(k)$
- > Clearly this event can happen... i.e. it is possible
- > But we have shown it has zero probability of occurring,  $P(x \in S(k)) = 0$
- > By the axioms of probability, the probability that it **doesn't happen** is,  $P(x \notin S(k)) = 1 - P(x \in S(k)) = 1$  We seem to have a paradox!

Solution Rethink how we interpret probability in continuous setting

- Define events as intervals instead of discrete values
- Assign probability to those intervals





Height represents *probability per unit* in the x-direction

We call this a **probability density** (as opposed to probability mass)

- > We denote the **probability density function** (PDF) as, p(X)
- > An event E corresponds to an *interval*  $a \le X < b$

> The probability of an interval is given by the area under the PDF,

$$P(a \le X < b) = \int_{a}^{b} p(X = x) \, dx$$

▷ Specific outcomes have zero probability  $P(X = 0) = P(x \le X < x) = 0$ 

> But may have nonzero *probability density* p(X = x)

**Definition** The <u>cumulative distribution function</u> (CDF) of a real-valued continuous RV X is the function given by,

 $F(x) = P(X \le x)$ 

Different ways to represent probability of interval, CDF is just a convention.

> Can easily measure probability of closed intervals,

$$P(a \le X < b) = F(b) - F(a)$$

 $\succ$  If X is absolutely continuous (i.e. differentiable) then,

$$f(x) = \frac{dF(x)}{dx}$$
 and  $F(t) = \int_{-\infty}^{t} f(x) dx$ 

Fundamental Theorem of Calculus

Where f(x) is the *probability density function* (PDF)

 $\blacktriangleright$  Typically use shorthand P for CDF and p for PDF instead of F and f

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Most definitions for discrete RVs hold, replacing PMF with PDF/CDF...

Two RVs X & Y are **independent** if and only if,

p(x,y) = p(x)p(y) or  $P(X \le x, Y \le y) = P(X \le x)P(Y \le y)$ 

**Conditionally independent** given Z iff,  $p(x, y \mid z) = p(x \mid z)p(y \mid z)$  or  $P(x, y \mid z) = P(x \mid z)P(y \mid z)$ 

Probability chain rule,

 $p(x,y) = p(x)p(y \mid x)$  and  $P(x,y) = P(x)P(y \mid x)$ 

...and by replacing summation with integration...

Law of Total Probability for continuous distributions,

$$p(x) = \int_{\mathcal{Y}} p(x, y) \, dy$$

Expectation of a continuous random variable,

$$\mathbf{E}[X] = \int_{\mathcal{X}} x \cdot p(x) \, dx$$

Covariance of two continuous random variables X & Y,

$$\mathbf{Cov}(X,Y) = \mathbf{E}[(X - \mathbf{E}[X])(Y - \mathbf{E}[Y])] = \int_{\mathcal{X}} \int_{\mathcal{Y}} (x - \mathbf{E}[X])(y - \mathbf{E}[Y])p(x,y) \, dx dy$$

**Caution** Some technical subtleties arise in continuous spaces...

For **discrete** RVs X & Y, the conditional

P(Y=y)=0 means impossible

$$P(X = x \mid Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

is **undefined** when  $P(Y=y) = 0 \dots$  no problem.

# For continuous RVs we have, $P(X \le x \mid Y = y) = \frac{P(X \le x, Y = y)}{P(Y = y)}$

but numerator and denominator are 0/0.

P(Y=y)=0 means improbable, but not impossible

Defining the conditional distribution as a limit fixes this...

$$\begin{split} P(X \leq x \mid Y = y) &= \lim_{\delta \to 0} P(X \leq x \mid y \leq Y \leq y + \delta) \\ &= \lim_{\delta \to 0} \frac{P(X \leq x, y \leq Y \leq y + \delta)}{P(y \leq Y \leq y + \delta)} \\ &= \lim_{\delta \to 0} \frac{P(X \leq x, Y \leq y + \delta) - P(X \leq x, Y \leq y)}{P(Y \leq y + \delta) - P(Y \leq y)} \\ &= \int_{-\infty}^{x} \lim_{\delta \to 0} \frac{\frac{\partial}{\partial x} P(u, y + \delta) - \frac{\partial}{\partial x} P(u, y)}{P(y + \delta) - P(y)} \, du \\ &= \int_{-\infty}^{x} \lim_{\delta \to 0} \frac{\left(\frac{\partial}{\partial x} P(u, y + \delta) - \frac{\partial}{\partial x} P(u, y)\right) / \delta}{(P(y + \delta) - P(y)) / \delta} \, du \\ &= \int_{-\infty}^{x} \frac{\frac{\partial^{2}}{\partial x \partial y} P(u, y)}{\frac{\partial}{\partial y} P(y)} \, du \quad = \int_{-\infty}^{x} \frac{p(u, y)}{p(y)} \, du \end{split}$$

Definition The conditional PDF is given by,  $p(x \mid y) = \frac{p(x,y)}{p(y)}$ 

(Fundamental theorem of calculus) (Assume interchange limit / integral)

( Multiply by  $rac{\delta}{\delta}=1$  )

( Definition of partial derivative ) ( Definition PDF )

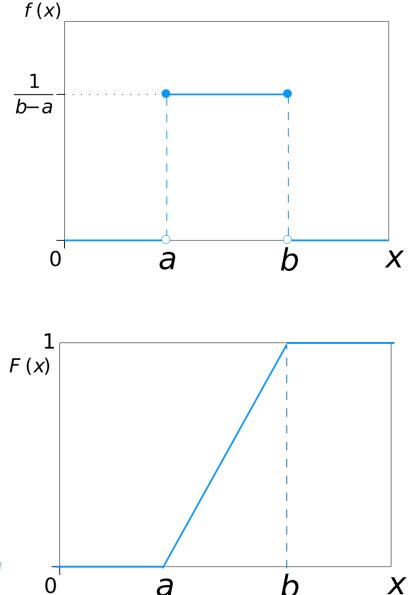
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**Uniform** distribution on interval [a, b],  $p(x) = \begin{cases} 0 & \text{if } x \le a, \\ \frac{1}{b-a} & \text{if } a \le x \le b, \\ 0 & \text{if } b < x \end{cases} \quad P(X \le x) = \begin{cases} 0 & \text{if } x \le a, \\ \frac{x-a}{b-a} & \text{if } a \le x \le b, \\ 1 & \text{if } b \le x \end{cases} \quad \frac{1}{b-a}$ Say that  $X \sim U(a, b)$  whose moments are,  $\mathbf{E}[X] = \frac{b+a}{2}$   $\mathbf{Var}[X] = \frac{(b-a)^2}{12}$ Suppose  $X \sim U(0,1)$  and we are told  $X \leq \frac{1}{2}$ what is the conditional distribution?  $P(X \le x \mid X \le \frac{1}{2}) = U(0, \frac{1}{2})$ 

Holds generally: Uniform closed under conditioning



**Gaussian** (a.k.a. Normal) distribution with mean mean (location)  $\mu$  and variance (scale)  $\sigma^2$  parameters,

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp{-\frac{1}{2}(x-\mu)^2/\sigma^2}$$

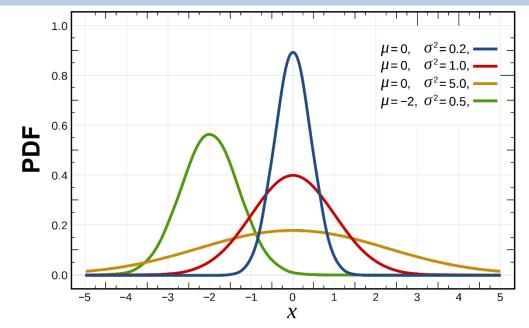
We say  $X \sim \mathcal{N}(\mu, \sigma^2)$  .

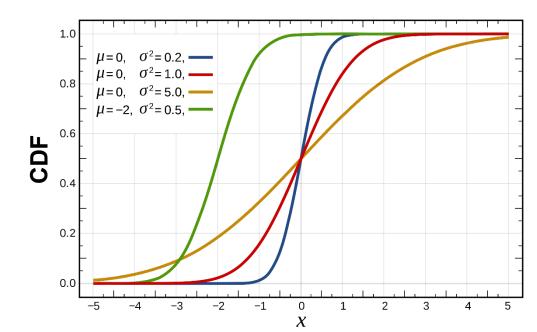
#### **Useful Properties**

• Closed under additivity:

 $X \sim \mathcal{N}(\mu_x, \sigma_x^2) \qquad Y \sim \mathcal{N}(\mu_y, \sigma_y^2)$  $X + Y \sim \mathcal{N}(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2)$ 

- Closed under linear functions (a and b constant):  $aX+b\sim \mathcal{N}(a\mu_x+b,a^2\sigma_x^2)$ 





Multivariate Gaussian On RV  $X \in \mathcal{R}^d$ with mean  $\mu \in \mathcal{R}^d$  and positive semidefinite covariance matrix  $\Sigma \in \mathcal{R}^{d \times d}$ ,

$$p(x) = |2\pi\Sigma|^{-1/2} \exp{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

Moments given by parameters directly.

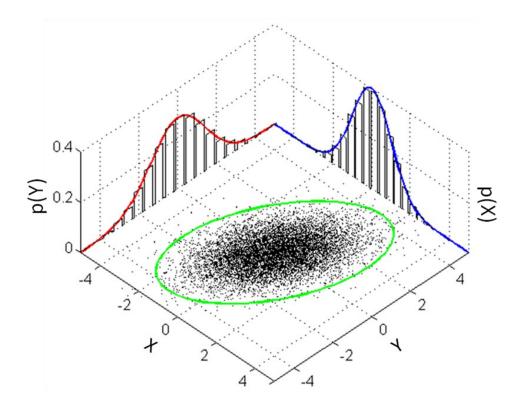
#### **Useful Properties**

- Closed under additivity (same as univariate case)
- Closed under linear functions,

 $AX + b \sim \mathcal{N}(A\mu_x + b, A\Sigma A^T)$ 

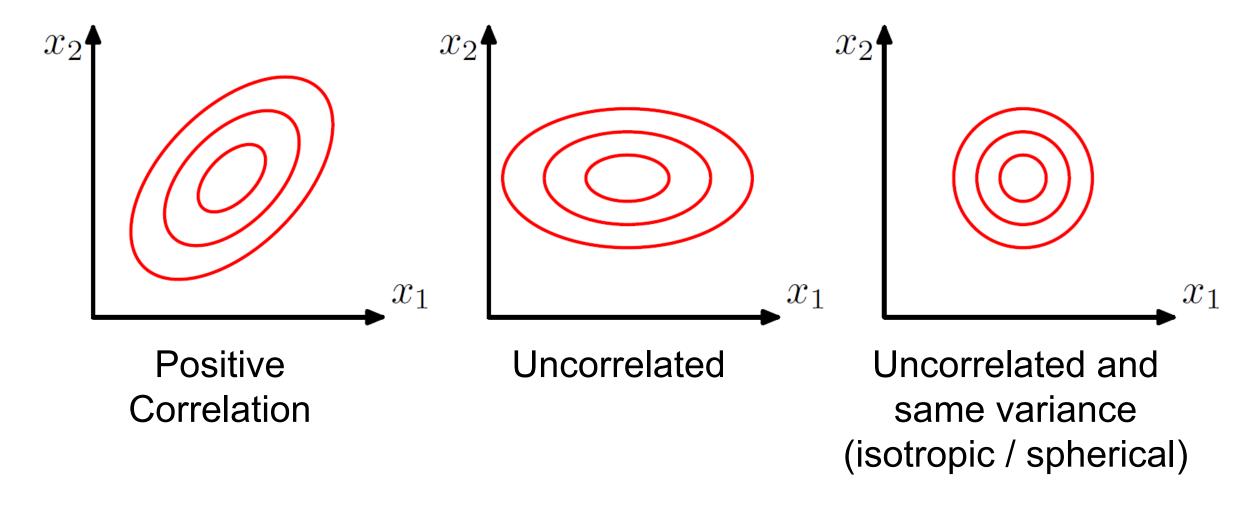
Where  $A \in \mathcal{R}^{m \times d}$  and  $b \in \mathcal{R}^m$  (output dimensions may change)

• Closed under conditioning and marginalization



#### Covariance

Captures correlation between random variables...can be viewed as set of ellipses...



#### **Covariance Matrix**

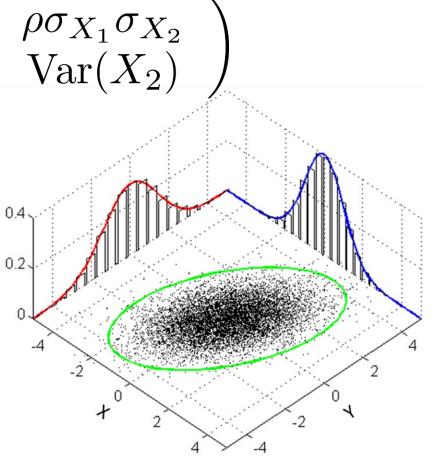
$$\Sigma = \operatorname{Cov}(X) = \begin{pmatrix} \operatorname{Var}(X_1) & \rho \sigma_{X_1} \sigma_{X_2} \\ \rho \sigma_{X_1} \sigma_{X_2} & \operatorname{Var}(X_2) \end{pmatrix}$$

#### **Covariance Matrix**

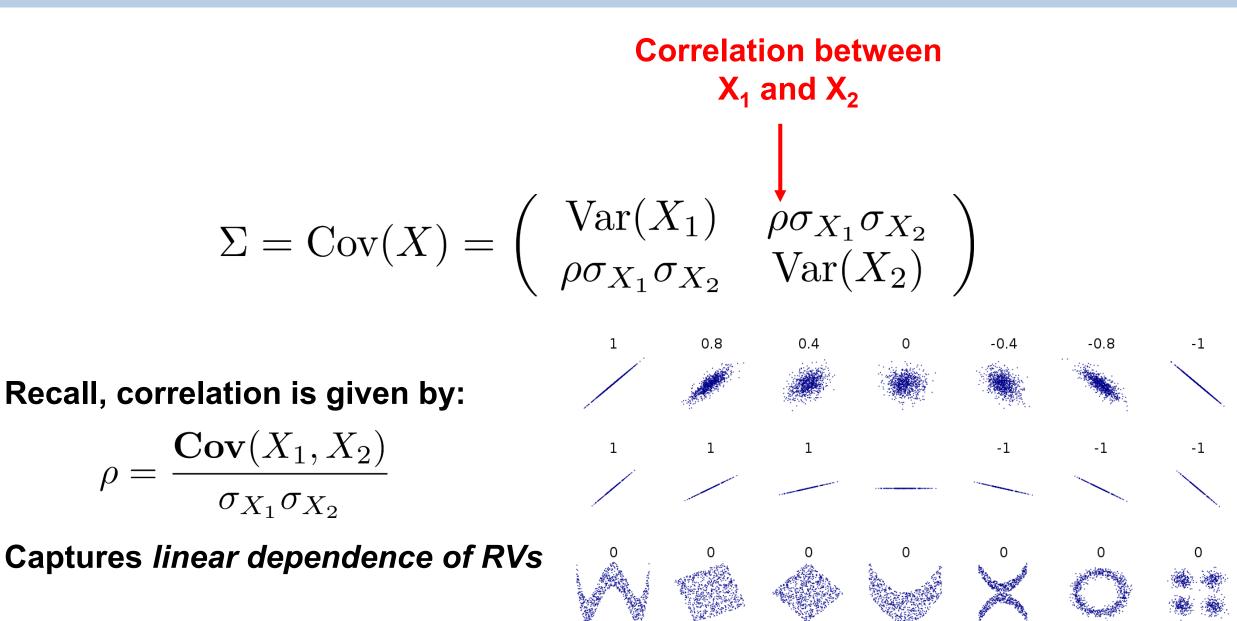
Marginal variance of just the RV X<sub>1</sub>

$$\Sigma = \operatorname{Cov}(X) = \begin{pmatrix} \operatorname{Var}(X_1) & \rho \\ \rho \sigma_{X_1} \sigma_{X_2} & \gamma \end{pmatrix}$$

i.e. How "spread out" is the distribution in the X<sub>1</sub> dimension...

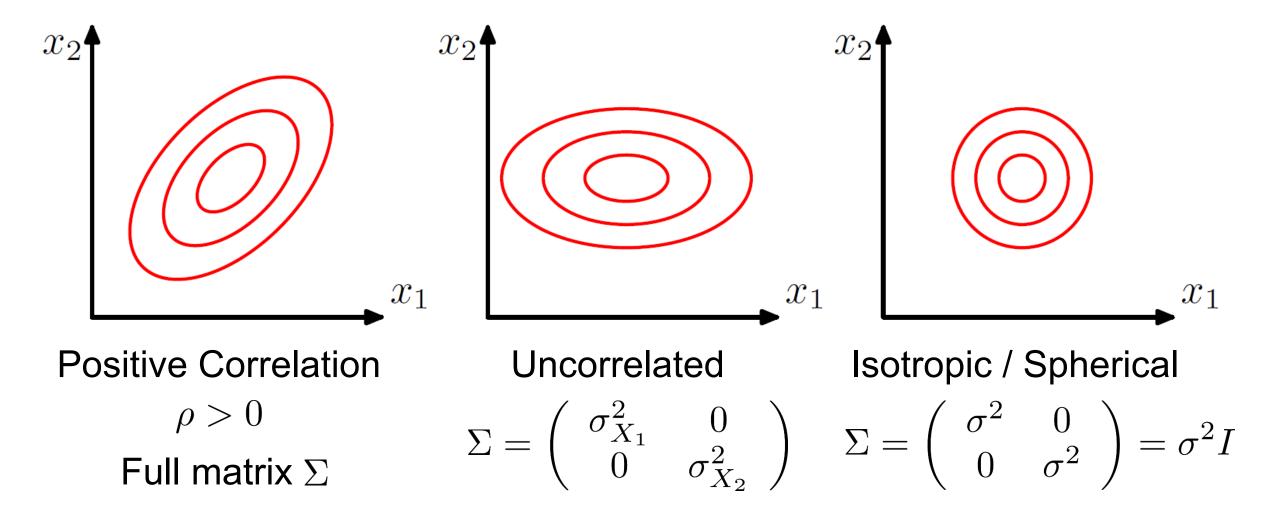


#### **Covariance Matrix**



### Covariance

Captures correlation between random variables...can be viewed as set of ellipses...



#### **Exponential** distribution with scale $\lambda$ ,

$$p(x) = \lambda e^{-\lambda x}$$
  $P(x) = 1 - e^{-\lambda x}$ 

for X>0. Moments given by,

$$\mathbf{E}[X] = \frac{1}{\lambda} \qquad \qquad \mathbf{Var}[X] = \frac{2}{\lambda^2}$$

#### **Useful properties**

• Closed under conditioning If  $X \sim \text{Exponential}(\lambda)$  then,

$$P(X \ge s + t \mid X \ge s) = P(X \ge s) = e^{-\lambda s}$$

• Minimum Let  $X_1, X_2, \ldots, X_N$  be i.i.d. exponentially distributed with scale parameters  $\lambda_1, \lambda_2, \ldots, \lambda_N$  then,

 $P(\min(X_1, X_2, \dots, X_N)) = \text{Exponential}(\sum_i \lambda_i)$ 

